

## Multi-product firm under uncertainty

### Setup

Consider a firm producing  $m$  products under:

$$\begin{array}{ll} \text{Output} & y = (y_1, \dots, y_m)^T \\ \text{Net return} & p = (p_1, \dots, p_m)^T \\ \text{Profit} & \pi = p^T y \\ \text{Utility function} & U(\pi) \end{array}$$

Notice that  $p$  denotes the net return, which means that  $p$  reflects both output prices and production costs per unit. Here, we assume only output price uncertainty, so only  $p$  are random variables (aka. a random vector). The problem is, then, to find

$$y^* = \operatorname{argmax}_{y \in \mathcal{Y}} \mathbb{E}[U(\pi)],$$

where  $\mathcal{Y}$  is a feasible set of production. To facilitate the analysis, we restrict our attention to the effects of correlation between  $p_i$  and  $p_j$  for all  $i, j \in \{1, \dots, m\}$  under the mean-variance model. Recall the mean-variance model (Chapter 6), which assumes

$$\mathbb{E}[U(\pi)] = W(\mathbb{E}[\pi], \operatorname{Var}[\pi]).$$

To find an optimal  $y^*$ , we need to write  $\mathbb{E}[\pi]$  and  $\operatorname{Var}[\pi]$  as functions of  $y$ .

$$\mathbb{E}[\pi] = \mathbb{E}[p^T y] = \mathbb{E}\left[\sum_{i=1}^m p_i y_i\right] = \sum_{i=1}^m \mathbb{E}[p_i] y_i = \sum_{i=1}^m \mu_i y_i = \mu^T y,$$

where  $\mu = (\mu_1, \dots, \mu_m)$  is the expected value of  $p$ , and

$$\operatorname{Var}[\pi] = \operatorname{Var}[p^T y] = y^T \operatorname{Var}[p] y = y^T \Sigma y,$$

where  $\Sigma$  is the covariance matrix of  $p$

$$\Sigma = \begin{bmatrix} \sigma_{11} & \cdots & \sigma_{1m} \\ \vdots & \ddots & \vdots \\ \sigma_{m1} & \cdots & \sigma_{mm} \end{bmatrix}, \quad \sigma_{ij} = \operatorname{Cov}(p_i, p_j).$$

Remember that  $U(\cdot)$  takes a scalar value  $\pi$ , so  $\mathbb{E}[\pi]$  and  $\operatorname{Var}[\pi]$  are scalar whereas  $p$  is a vector and  $\Sigma$  is a matrix. Having written  $\mathbb{E}[\pi]$  and  $\operatorname{Var}[\pi]$  as functions of  $y$ , the problem is now to find

$$y^* = \operatorname{argmax}_{y \in \mathcal{Y}} W(\mathbb{E}[\pi], \operatorname{Var}[\pi]). \quad (1)$$

**E-V frontier**

If the form of expected utility  $W(\cdot)$  is known, we may directly solve (1). If not, we may take a different approach. Recall that under the mean-variance model with risk aversion assumed, for given  $M = \mu^T y$ , it is optimal to minimize  $\text{Var}[\pi]$ .

$$\begin{aligned} y^+(M) &= \underset{y \in \mathcal{Y}}{\text{argmin}} \text{Var}[\pi] \text{ s.t. } M = \mu^T y \\ &= \underset{y \in \mathcal{Y}}{\text{argmin}} y^T \Sigma y \text{ s.t. } M = \mu^T y. \end{aligned} \quad (2)$$

This is a quadratic programming (Boyd & Vandenberghe, 2004), i.e. a quadratic objective function with linear constraints. So, there exists a solution. Indeed, we may set up the Lagrangian:

$$\mathcal{L}(y, \lambda) = y^T \Sigma y - \lambda(\mu^T y - M)$$

and examine the first order conditions

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial y} &= 2\Sigma y - \lambda\mu = 0 \\ \mu^T y &= M \end{aligned}$$

or

$$\begin{pmatrix} 2\Sigma & -\mu \\ \mu^T & 0 \end{pmatrix} \begin{pmatrix} y \\ \lambda \end{pmatrix} = \begin{pmatrix} 0 \\ M \end{pmatrix} \quad (3)$$

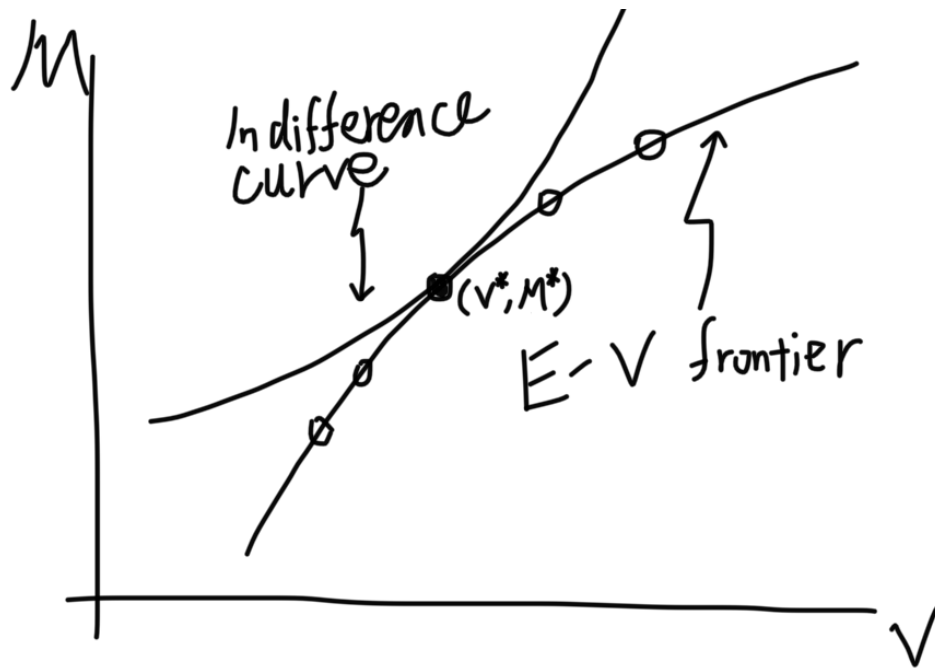
Note that the matrix on the left-hand side is not necessarily invertible.

The E-V frontier is a set

$$\left\{ (V, M) \in \mathbb{R}_+^2 : V = \min_{y \in \mathcal{Y}} y^T \Sigma y \text{ s.t. } M = \mu^T y \right\}.$$

We can also see it as a graph and plot it with  $V$  on the horizontal axis and  $M$  on the vertical axis.

With the E-V frontier in hand, we may ask a firm manager which point  $(V, M)$  on the curve he/she prefers. The point chosen  $(V^*, M^*)$  tells us a solution  $y^* = y^+(M^*)$ . Even without the knowledge of  $W(\cdot)$  or indifference curves, picking  $(V^*, M^*)$  implies that the unknown indifference curve is tangent at that point.



### Diversification

Let's look at the effects of correlation under  $m = 2$ . Let  $\rho$  be the correlation between  $p_1$  and  $p_2$ . Then, rewrite the covariance matrix as follows

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix}.$$

Notice, for brevity,  $\sigma_1^2 = \text{Var}[p_1] = \sigma_{11}$  and  $\sigma_2^2 = \text{Var}[p_2] = \sigma_{22}$  in the previous notation. Observe,

$$\begin{aligned} \text{Var}[\pi] &= y^T \Sigma y \\ &= \sigma_1^2 y_1^2 + \sigma_2^2 y_2^2 + 2\rho\sigma_1\sigma_2 y_1 y_2. \end{aligned}$$

Clearly, if  $\rho = -1$ ,

$$\begin{aligned} \text{Var}[\pi] &= \sigma_1^2 y_1^2 + \sigma_2^2 y_2^2 - 2\sigma_1\sigma_2 y_1 y_2 \\ &= (\sigma_1 y_1 - \sigma_2 y_2)^2, \end{aligned}$$

which is minimized to 0 by choosing

$$\begin{aligned} \sigma_1 y_1 &= \sigma_2 y_2 \\ \Leftrightarrow \sigma_1 y_1 &= \sigma_2 \frac{M - \mu_1 y_1}{\mu_2} \\ \Leftrightarrow y_1 &= \frac{M\sigma_2}{\mu_2\sigma_1 + \mu_1\sigma_2}. \end{aligned}$$

Hence, with the perfect negative correlation ( $\rho = -1$ ), we can completely eliminate the risk regardless of  $M$ . That is, the E-V frontier is a vertical line at  $V = 0$ .

How about  $\rho = 0$ ? Notice that, with  $\rho = 0$ ,  $\Sigma$  is diagonal and positive definite. So, the matrix in (3) is invertible.

$$\begin{aligned} \begin{pmatrix} y \\ \lambda \end{pmatrix} &= \begin{pmatrix} 2\Sigma & -\mu \\ \mu^T & 0 \end{pmatrix}^{-1} \begin{pmatrix} 0 \\ M \end{pmatrix} \\ &= \frac{1}{\mu_1^2\sigma_2^2 + \mu_2^2\sigma_1^2} \begin{pmatrix} \frac{1}{2}\mu_2^2 & -\frac{1}{2}\mu_1\mu_2 & \mu_1\sigma_2^2 \\ -\frac{1}{2}\mu_1\mu_2 & -\frac{1}{2}\sigma_1^2 & \mu_2\sigma_1^2 \\ -\mu_1\sigma_2^2 & -\mu_2\sigma_1^2 & 2\sigma_1^2\sigma_2^2 \end{pmatrix} \begin{pmatrix} 0 \\ M \end{pmatrix}, \end{aligned}$$

which implies

$$\begin{pmatrix} y_1 \\ y_2 \end{pmatrix} = \frac{M}{\mu_1^2\sigma_2^2 + \mu_2^2\sigma_1^2} \begin{pmatrix} \mu_1\sigma_2^2 \\ \mu_2\sigma_1^2 \end{pmatrix}.$$

As a result,

$$\begin{aligned} V &= \text{Var}[\pi] \\ &= y^T \Sigma y \\ &= \begin{pmatrix} \sigma_1^2 & \sigma_2^2 \end{pmatrix} \begin{pmatrix} y_1^2 \\ y_2^2 \end{pmatrix} \\ &= \begin{pmatrix} \sigma_1^2 & \sigma_2^2 \end{pmatrix} \left( \frac{M}{\mu_1^2\sigma_2^2 + \mu_2^2\sigma_1^2} \right)^2 \begin{pmatrix} \mu_1^2\sigma_2^4 \\ \mu_2^2\sigma_1^4 \end{pmatrix} \\ &= \frac{\mu_1^2\sigma_1^2\sigma_2^4 + \mu_2^2\sigma_1^4\sigma_2^2}{(\mu_1^2\sigma_2^2 + \mu_2^2\sigma_1^2)^2} M^2. \end{aligned}$$

Hence, the E-V frontier is  $M = \alpha\sqrt{V}$  where  $\alpha > 0$ , which means that it is an upward straight line if  $M$  is plotted against standard deviation.